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## A NOTE ON LIMITING BEHAVIOUR OF DISASTROUS ENVIRONMENT EXPONENTS

**Thomas S. Mountford**

Department of Mathematics  
University of California, Los Angeles  
Los Angeles, CA 90095–1555  
[malloy@math.ucla.edu](mailto:malloy@math.ucla.edu)

**Abstract** We consider a random walk on the  $d$ -dimensional lattice and investigate the asymptotic probability of the walk avoiding a "disaster" (points put down according to a regular Poisson process on space-time). We show that, given the Poisson process points, almost surely, the chance of surviving to time  $t$  is like  $e^{-\alpha \log(\frac{1}{k})t}$ , as  $t$  tends to infinity if  $k$ , the jump rate of the random walk, is small.

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## Introduction

This note concerns a recent work of T. Shiga ([**Shi**]). The following model was considered: We are given a system of independent rate one Poisson processes on  $[0, \infty)$ ,  $\underline{N} = \{N_x(t)\}_{x \in \mathbb{Z}^d}$ . We are also given an independent simple random walk on  $\mathbb{Z}^d$ ,  $X(t)$ , moving at rate  $k$  and with, say,  $X(0) = \underline{0}$ .

Of course simply by integrating out over  $\underline{N}, X$  we have (taking  $\delta N_{X(s)}(s) = N_{X(s)}(s) - N_{X(s)}(s-)$ )

$$\forall t \geq 0 \quad P[\forall 0 \leq s \leq t \delta N_{X(s)}(s) = 0] = e^{-t}.$$

The problem becomes non-trivial when considering

$$\begin{aligned} p(t, N) &= P[\forall 0 \leq s \leq t \delta N_{X(s)}(s) = 0 | \underline{N}] = \\ &P[\forall 0 \leq s \leq t \delta N_{X(s)}(s) = 0 | \underline{N}(s) \ s \leq t]. \end{aligned}$$

It is non-trivial, but was shown in [**Shi**], that the random quantity  $p(t, N)$  satisfies

$$\lim_{t \rightarrow \infty} \frac{\log p(t, N)}{t} = -\lambda(d, k)$$

It was shown that as  $k$  becomes large  $\lambda$  tends to one in all dimensions and that in dimensions three and higher  $\lambda$  is equal to one for  $k$  sufficiently large. The focus of this note is on the other behaviour of  $\lambda(d, k)$ : the behaviour as  $k \rightarrow 0$ . It was shown in [**Shi**] that there existed two constants  $c_1, c_2 \in (0, \infty)$  so that

$$c_1 < \liminf_{k \rightarrow 0} \frac{\lambda(d, k)}{\log(\frac{1}{k})} \leq \limsup_{k \rightarrow 0} \frac{\lambda(d, k)}{\log(\frac{1}{k})} < c_2.$$

We wish to show

**Theorem 1.0** *There exists a constant  $\alpha$  so that  $\lim_{k \rightarrow 0} \frac{\lambda(d, k)}{\log(\frac{1}{k})} = \alpha$ .*

The paper is organized as follows: in Section One we consider a "shortest path" problem which is easily and naturally dealt with by Liggett's subadditive ergodic theorem (see [**L**]). This yield a constant  $\alpha$ . In Section Two we show (Corollary 2.4) that  $\liminf_{k \rightarrow 0} \frac{\lambda(d, k)}{\log(\frac{1}{k})} \geq \alpha$  and in Section Three we show (Corollary 3.1)  $\limsup_{k \rightarrow 0} \frac{\lambda(d, k)}{\log(\frac{1}{k})} \leq \alpha$ , thus completing the proof of Theorem 1.0.

Both of the last two sections rely heavily on block arguments as popularized in [**D**], [**D1**].

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## Section One

In this section we consider only the Poisson processes  $N$ . The random walk will not be directly considered at all, though sometimes it will be implicit, as in the definition of a *path* below:

A path  $\gamma$  is a piecewise constant right continuous function with left limits  $[0, \infty) \rightarrow \mathbb{Z}^d$  so that for all  $t$   $\|\gamma(t) - \gamma(t-)\|_1 \leq 1$ .

The collection of paths beginning at  $x \in \mathbb{Z}^d$  which avoid points in  $N$  up to time  $t$  will be denoted by  $\Gamma^{x,t}$ . More formally

$$\Gamma^{x,t} = \{ \gamma : \forall 0 \leq s \leq t \delta N_{\gamma(s)}(s) = 0, \gamma(0) = x \}.$$

(Again, consistent with previous notation,  $\delta N_{\gamma(s)} = N_{\gamma(s)}(s) - N_{\gamma(s)}(s-)$ .) For  $\gamma \in \Gamma^{x,t}$ ,  $S^x(\gamma, t) = \sum_{0 \leq s \leq t} I_{\gamma(s) \neq \gamma(s-)}$  where  $I$  is the usual indicator function. In words  $S$  counts the number of jumps that  $\gamma$  makes in time interval  $[0, t]$ . If  $x = 0$  we suppress the suffix  $x$ .

Finally we define

$$\alpha(t, N) = \min\{S(\gamma, t) : \gamma \in \Gamma^t = \Gamma^{0,t}\}.$$

**Proposition 1.1**  $\alpha = \lim_{t \rightarrow \infty} \frac{1}{t} \alpha(t, N)$  exists.

*Proof* Define random variables  $X_{s,t}$  for  $0 \leq s < t < \infty$  by

$$X_{0,t} = \alpha(t, N)$$

and for  $0 < s < t$

$$X_{s,t} = \inf\{S(\gamma, t) - S(\gamma, s) : \gamma \in \Gamma^t, \gamma(s) = x_s\}$$

where  $x_s = \min\{x \in \mathbb{Z}^d : \exists \gamma \in \Gamma^s \text{ so that } S(\gamma, s) = \alpha(s, N), \gamma(s) = x\}$  under any well ordering of the points  $x \in \mathbb{Z}^d$ .

Then the random variables satisfy the conditions for Liggett's subadditive ergodic theorem. Given the ergodicity of our Poisson processes we conclude that the a.s. limit of  $\frac{1}{t} \alpha(t, N)$  is non random. □

We now show that the constant  $\alpha$  of Proposition 1.1 is strictly positive. This fact will follow from Theorem 1.0 and the results of [Shi], however we include it for completeness and because the argument given is a precursor to the block argument of Proposition 2.2.

**Proposition 1.2** *The constant  $\alpha$  is strictly positive.*

Fix  $\varepsilon > 0$  small we shall give conditions on the smallness of  $\varepsilon$  as the proof progresses. Choose integer  $L$  so that  $L^d e^{-L} < \varepsilon$ .

We divide up space time into cubes

$$V(\underline{n}, r) = [n_1 L, (n_1 + 1)L) \times [n_2 L, (n_2 + 1)L) \times \cdots \times [n_d L, (n_d + 1)L) \times [rL, (r + 1)L).$$

We associate 0-1 random variables  $\psi(\underline{n}, r)$  to these cubes by taking  $\psi(\underline{n}, r)$  to be 1 if and only if

$$\forall x \in [n_1 L, (n_1 + 1)L) \times [n_2 L, (n_2 + 1)L) \times \cdots \times [n_d L, (n_d + 1)L)$$

$$N_x((r+1)L-) - N_x(rL) \geq 1.$$

We note that the  $\psi$  random variables are *i.i.d.* and that, by the choice of  $L$ , the probability that  $\psi(\underline{n}, r) \neq 1$  is  $< \varepsilon$ .

To show our result it is sufficient to show that as  $m$  tends to infinity  $\alpha(mL, N) \geq \frac{m}{2}$  with probability tending to one.

The *trace* of a path  $\gamma \in \Gamma^{mL}$  is the sequence of points in  $\mathbb{Z}^d$ ,  $\underline{n}_i$   $0 \leq i \leq m$  so that for  $0 \leq i \leq m$ ,

$$(\gamma(iL), iL) \in V(\underline{n}_i, iL).$$

The crucial observation is that for such  $\gamma, \underline{n}_i$ ,

$$S^0(\gamma, mL) \geq \sum_{i=0}^{m-1} \psi(\underline{n}_i, i) + L \sum_{i=0}^{m-1} (\|\underline{n}_{i+1} - \underline{n}_i\|_\infty - 1)_+$$

since if  $\psi(\underline{n}_i, i) = 1$  then  $\gamma$  must make at least one jump in the time interval  $[iR, (i+1)R)$  and if, furthermore  $(\|\underline{n}_{i+1} - \underline{n}_i\|_\infty - 1)_+ = f$ , then in this time interval  $\gamma$  must make more than  $fL$  jumps.

Thus to show that  $\alpha(mL, N) \geq \frac{m}{2}$  it suffices to show that for all  $\{\underline{n}_i\}$  with

$$\sum_{i=0}^{m-1} (\|\underline{n}_{i+1} - \underline{n}_i\|_\infty - 1)_+ \leq \frac{m}{2L} \quad (1)$$

it is the case that

$$\left\{ \sum_{i=0}^{m-1} \psi(\underline{n}_i, i) \geq \frac{m}{2} \right\}.$$

By simple large deviations arguments the probability that for any given  $\{\underline{n}_i\}$ ,  $\{\sum_{i=0}^{m-1} \psi(\underline{n}_i, i) \geq \frac{m}{2}\}$  is less than  $2^m(\varepsilon)^{\frac{m}{2}}$ . Thus it remains only to count the number of  $\{\underline{n}_i\}$  satisfying (1).

We write (for positive integer  $g_i$ )  $A(g_1, g_2, \dots, g_m)$  for the set of  $(\underline{n}_1, \underline{n}_2, \dots, \underline{n}_m)$  so that for  $1 \leq i \leq m$ ,  $(\|\underline{n}_i - \underline{n}_{i-1}\|_\infty - 1)_+ = g_i$ . We first give a crude bound on the cardinality of  $A(g_1, g_2, \dots, g_m)$ :  $\underline{n}_0$  is required to be  $\underline{0}$ , after having "chosen"  $\underline{n}_0, \underline{n}_1 \dots \underline{n}_{i-1}$  we have  $3^d$  choices for  $\underline{n}_i$  if  $g_i = 0$ , otherwise we have at most  $2d(2g_i + 3)^{d-1}$  choices for  $\underline{n}_i$ . Thus (using  $2d \leq 3^d$ )

$$|A(g_1, g_2, \dots, g_m)| \leq 3^{md} \prod (2g_i + 3)^{d-1}.$$

We may find  $K$  so that for all  $g$ ,  $(2g + 3)^{d-1} \leq K2^g$ ; we conclude that

$$|A(g_1, g_2, \dots, g_m)| \leq 3^{md} K^{m2} 2^{\sum_{i=1}^m g_i} \leq C^m$$

for some universal  $C$  not depending on  $d, \varepsilon$ , if  $\sum g_i \leq \frac{m}{2L}$ .

By elementary combinatorics the number of  $(g_1, g_2, \dots, g_m)$  so that  $\sum_{i=1}^m g_i = r$  is  $\binom{m+r-1}{r}$ , thus the number of  $(g_1, g_2, \dots, g_m)$  so that  $\sum_{i=1}^m g_i \leq \frac{m}{2L}$  is less than  $2^{2m}$ . We conclude that the number of  $\{\underline{n}_i\}$  satisfying (1) is bounded by  $(4C)^m$ . Thus the probability that  $\alpha(mL, N)$  exceeds  $\frac{m}{2}$  is at least  $1 - (4C)^m 2^m (\varepsilon)^{\frac{m}{2}}$ . This tends to one as  $m$  tends to infinity provided that  $\varepsilon$  was fixed sufficiently small.

□

## Section Two

Fix  $\varepsilon > 0$ , arbitrarily small. Given  $c > 0$  fixed, we say that a cube  $[-cR, cR]^d$  is *good* if  $\forall x \in [-cR, cR]^d$

$$\inf_{\gamma \in \Gamma^{x,R}} S^x(\gamma, R) \geq R(\alpha - \varepsilon).$$

**Lemma 2.1** *Given  $\delta, c > 0$ , there exists  $R_0 = R_0(c, \delta)$  so that for all  $R \geq R_0$ ,*

$$P \left[ [-cR, cR]^d \text{ is good} \right] \geq 1 - \delta$$

*Proof* Given  $\varepsilon, c$ , there exists  $k$  so that for any  $R$ , we can pick points  $x_1^R, x_2^R \dots x_{k/\varepsilon^d}^R \in [-cR, cR]^d$  so that every point of  $[-cR, cR]^d$  is within  $R\varepsilon/10$  of  $x_j^R$  for at least one  $j$ . Given this property it is clear that event

$$\left\{ \inf_{x \in [-cR, cR]^d} \inf_{\gamma \in \Gamma^{x,R}} S(\gamma, R) < R(\alpha - \varepsilon) \right\}$$

is contained in

$$\left\{ \inf_{x_j^R} \inf_{\gamma \in \Gamma^{x_j^R, R}} S(\gamma, R) < R(\alpha - \varepsilon/2) \right\}$$

Thus we have

$$P \left[ [-cR, cR]^d \text{ is good} \right] \geq 1 - \frac{k}{\varepsilon^d} P[\alpha(R, N) < R(\alpha - \varepsilon/2)].$$

This last term is greater than  $1 - \delta$  if  $R$  is sufficiently large. □

We have not fully specified how small we require  $\delta$  to be but, conditional on this we will fix  $R$  at a level so large that the conclusions of Lemma 2.1 hold for  $\delta$  and also so that  $\gg \frac{1}{\varepsilon}$ .

**Lemma 2.2** *Given  $c$  and  $R \geq R_0$  fixed, there exists  $k_0 > 0$  so that if  $0 < k \leq k_0$  and cube  $[-cR, cR]^d$  is good then for any random walk  $X(t)$  starting in the cube, the chance of survival to time  $R$  is bounded above by  $k^{R(\alpha-2\varepsilon)}$ . More generally given  $c$ ,  $R \geq R_0$  we have for  $k \leq k_0$  that the chance that the random walk makes  $\geq f\alpha R$  jumps in time  $R$  is bounded above by  $k^{fR(\alpha-\varepsilon)}$ .*

*Proof* Let the starting point of  $X$  be  $x$ . By definition of  $\alpha(R, N)$  and a cube being good we have

$$P[X(\cdot) \in \Gamma^{x,R}] \leq P[S(X(\cdot), R) \geq R(\alpha - \varepsilon)] \leq (Rk)^{R(\alpha-\varepsilon)}.$$

This latter term is less than  $k^{R(\alpha-2\varepsilon)}$  if  $k$  is sufficiently small. □

We choose  $c$  to equal  $10(\alpha + 1)$  and divide up the lattice into cubes  $C(\underline{n}) = 2cR\underline{n} + [-cR, cR]^d$ . We divide up space time into cubes  $D(\underline{n}, i) = C(\underline{n}) \times [iR, (i+1)R]$ . We say that  $D(\underline{n}, i)$  is good if  $[-cR, cR]^d$  is good (in the old sense) after translating Poisson system  $(\underline{N})$  spatially by  $2cR\underline{n}$  and temporally by  $iR$ .

We define random variables  $\psi(\underline{n}, i)$  taking values 0 or 1 by

$$\psi(\underline{n}, i) = 1 \text{ if } D(\underline{n}, i) \text{ is good.}$$

The random variables  $\psi(\underline{n}, i)$  are not independent, but it should be noted that random variables  $\psi(\underline{n}_1, i_1), \psi(\underline{n}_2, i_2) \cdots \psi(\underline{n}_j, i_j)$  are independent if the  $i_h$  s are all distinct.

A  $v$ -chain  $\beta$  is a sequence  $(\beta_j, j) \ j = 0, 1, \dots, v-1$ . We do not require that  $|\beta_{j+1} - \beta_j|_1$  be less than or equal to 1.

An  $(r-v)$ -chain is a sequence  $(\beta_j, j) \ j = r, r+1, \dots, v-1$ .

Given  $\psi$  we associate a score to a  $(r-v)$ -chain  $\beta$  by

$$J_v(\beta) = \sum_{j=r}^{j=v-1} \psi(\beta_j, j) + 9 \sum_{j=r}^{j=v-2} (|\beta_{j+1} - \beta_j|_\infty - 1)_+.$$

**Proposition 2.1** *For a random walk starting at time  $rR$  in cube  $C(\underline{n})$ , the chance that it survives until time  $vR$  is bounded above by*

$$2^{v-r-1} \exp \left( R(\alpha - 2\varepsilon) \ln(k) \min_{\beta} J_v(\beta) \right)$$

where the minimum is taken over all  $(r-v)$ -chains  $\beta$  with  $\beta_r = \underline{n}$ .

*Proof* In the proof we regard  $v$  as fixed and use induction on  $k = v - r$ . The proof follows from induction on  $k$ . It is clearly true for  $k = 1$  (or  $r = v - 1$ ) and all  $\underline{n}$  by Lemma 2.2. Suppose that it is true for  $k - 1$  (and all possible  $\underline{n}$ ) and suppose further that  $X^k$  is a random walk starting at time  $R(v - k)$  in cube  $C(\underline{n})$ . We consider the random walk over time interval  $[(v - k)R, (v - k + 1)R]$ .

$$\begin{aligned} P[X^k \text{ survives up to } vR] &= \\ \sum_{\underline{m}} P[X^k \text{ survives up to } (v - k + 1)R, \\ X^k((v - k + 1)R) \in C(\underline{m}), X^k \text{ survives up to } vR]. \end{aligned}$$

By the Markov property for  $X^k$  and induction this summation is bounded by

$$\begin{aligned} \sum_{\underline{m}} P[X^k \text{ survives up to } (v - k + 1)R, \\ X^k((v - k + 1)R) \in C(\underline{m})] (2^{k-2}) \exp(R(\alpha - 2\varepsilon) \ln(k) J_v^{\underline{m}, k-1, v}) \end{aligned}$$

where  $J_v^{\underline{m}, k-1, v}$  is the minimum of  $J_v(\beta)$  over  $(v - k + 1) - v$ -chains  $\beta$  with  $\beta_{v-k+1} = \underline{m}$ . This in turn is majorized by

$$\sum_{f=2} P[X^k \text{ survives up to } (v - k + 1)R, X^k((v - k + 1)R) \in C(\underline{m})]$$

$$\begin{aligned}
& \text{with } \|\underline{n} - \underline{m}\|_\infty = f] (2^{k-2}) \exp(R(\alpha - 2\varepsilon) \ln(k) J_v^{f,k-1,v}) \\
& + \sum_{\|\underline{n} - \underline{n}'\|_\infty \leq 1} P[X^r \text{ survives up to } (v - k + 1)R, X^k((v - k + 1)R) \in C(\underline{n}')] \\
& \quad (2^{k-2}) \exp(R(\alpha - 2\varepsilon) \ln(k) J_v^{\underline{n}',k-1,v})
\end{aligned}$$

for  $J_v^{f,k-1,v}$  the minimum of  $J_v^{m,k-1,v}$  over  $\|\underline{n} - \underline{m}\|_\infty = f$ . By Lemma 2.2 these two summations are bounded by

$$\begin{aligned}
& (2^{k-2}) \exp\left((\alpha - 2\varepsilon)R \ln(k)(\psi(\underline{n}, r) + J_v^{1,k-1,v})\right) + \\
& (2^{k-2}) \sum_{f=2}^{\infty} \exp\left((f-1)10\alpha R \ln(k) + R(\alpha - 2\varepsilon) \ln(k) J_v^{f,k-1,v}\right)
\end{aligned}$$

where  $J_v^{1,k-1,v}$  is the minimum of  $J_v^{m,k-1,v}$  over  $\|\underline{n} - \underline{m}\|_\infty \leq 1$  (a slightly different definition from that of  $J_v^{f,k-1,v}$  for higher  $f$ ).

If  $R$  was chosen sufficiently large this is bounded by

$$2^{k-1} \exp\left(R(\alpha - 2\varepsilon) \ln(k) \min_{\beta} J_v(\beta)\right)$$

where the minimum is taken over all (r-v)-chains  $\beta$  with  $\beta_r = \underline{n}$ .  $\square$

It remains to show that as  $v$  tends to infinity  $J_v(\beta)$  is roughly  $v$ . It is time to properly define  $\delta$ . First fix  $K \gg 3^d$  and so that for each integer  $f$  at least 1, the number of  $\underline{m}$  with  $\|\underline{m}\|_\infty = f$  is less than  $K2^{f-1}/100$ .

**Lemma 2.3** *Given  $\varepsilon > 0$  there exists  $\delta$  so that  $0 < \delta < \varepsilon/100K$  so that if  $X_1, X_2, \dots, X_N$  are i.i.d. Bernoulli  $\delta$  random variables for any integer  $N$  then*

$$P\left[\sum_{j=1}^N X_j \geq N\varepsilon + r\right] \leq \left(\frac{1}{100K}\right)^{N+r}.$$

**Proposition 2.2** *With probability one for all  $v$  sufficiently large*

$$\inf_{\beta \in J_v} J(\beta) \geq v(1 - 2\varepsilon)$$

*Proof* We simply count. Given our definition of  $J(\beta)$  we need only consider those  $\beta \in J_v$  with  $\sum_{j=0}^{v-2} (\|\beta_{j+1} - \beta_j\|_\infty - 1)_+ \leq v/9$ . For  $\beta \in J_v$ , we say the code of  $\beta$  is the sequence

$$\{(\|\beta_1 - \beta_0\|_\infty - 1)_+ \cdots (\|\beta_{j+1} - \beta_j\|_\infty - 1)_+ \cdots (\|\beta_{v-1} - \beta_{v-2}\|_\infty - 1)_+\}.$$

For fixed code  $m_0, m_1 \cdots m_{v-2}$  with  $\sum m_j \leq v/9$  there are (by our choice of  $K$ ) less than or equal to  $K^{v-1} \prod_{j=0}^{v-2} 2^{m_i-1}$  possible v-chains. For any such  $\beta$ ,  $J_v(\beta) = 9 \sum m_j + \sum \psi(\beta_j, j)$  and so

$$P[J_v(\beta) \leq v(1 - 2\varepsilon)] \leq P\left[\sum \psi(\beta_j, j) \leq v(1 - 2\varepsilon) - 9 \sum m_j\right]$$

$$= P[\sum (1 - \psi(\beta_j, j)) \geq v2\varepsilon + 9 \sum m_j] \leq \left(\frac{1}{100K}\right)^v \left(\frac{1}{100K}\right)^{9 \sum m_j}.$$

So the probability that for some  $\beta$  with code  $m_0, m_1, \dots, m_{v-2}$   $J_v(\beta)$  is less than or equal to  $(1 - 2\varepsilon)v$  is bounded by

$$K^v \prod_{j=0}^{j=v-2} 2^{m_i-1} \left(\frac{1}{100K}\right)^v \left(\frac{1}{100K}\right)^{\sum m_j} \leq \left(\frac{1}{100}\right)^v.$$

But the number of codes which sum to less than  $v/9$  is (assuming w.l.o.g that  $v/9$  is an integer) exactly  $\sum_{j=0}^{j=v/9} \binom{v+j-1}{v-1} \leq v/9 \binom{v+v/9-1}{v-1} \leq 2^v$  for  $v$  large. We conclude that  $P[\min J_v(\beta) \leq v(1 - 2\varepsilon)] \leq \left(\frac{1}{50}\right)^v$  for large  $v$ . The proposition now follows from the Borel Cantelli Lemma.  $\square$

**Corollary 2.4**  $\liminf_{k \rightarrow 0} \frac{\lambda(d, k)}{\ln(\frac{1}{k})} \geq \alpha(d)$ .

*Proof* By Proposition 2.1 we have that for  $k \leq k_0$  that

$$p(vR, N) \leq 2^{v-1} \exp(R(\alpha - 2\varepsilon) \ln(k) \min J_v(\beta))$$

By Proposition 2.2 we have therefore that for large enough  $v$

$$\begin{aligned} p(vR, N) &\leq 2^{v-1} \exp(R(\alpha - 2\varepsilon) \ln(k) v(1 - 2\varepsilon)) \\ &\leq 2^{vR\varepsilon} \exp(R(\alpha - 2\varepsilon) \ln(k) v(1 - 2\varepsilon)) \\ &\leq \exp(Rv((\alpha - 2\varepsilon) \ln(k)(1 - 2\varepsilon) + \varepsilon)) \end{aligned}$$

Thus we have that  $\lambda(k, d) \geq \ln(\frac{1}{k})(\alpha - 2\varepsilon)(1 - 2\varepsilon) - \varepsilon$ . Since  $\varepsilon$  is arbitrarily small the Corollary follows.  $\square$

## Section Three

In this section we will use block/percolation arguments that since [BG] may be regarded as standard. Simply to avoid notational encumbrance we will write out the proof for the case  $d = 1$  but the argument easily extends to all dimensions.

Fix  $\varepsilon > 0$ . By Proposition 1.1 we have that for  $R$  sufficiently large

$$P[\alpha(R, N) \leq R(\alpha + \varepsilon)] > 1 - \varepsilon^6.$$

Now note that, by our definition of  $\alpha$ , the event  $\{\alpha(R, N) \leq R(\alpha + \varepsilon)\}$  is the same as the event  $\{\exists \gamma \in \Gamma^R$  with  $S(\gamma, R) \leq R(\alpha + \varepsilon)$  and  $|\gamma(R)| \leq R(\alpha + \varepsilon)\}$ . Thus for  $R$  sufficiently large

$$\begin{aligned} &\{\nexists \gamma \in \Gamma^R \text{ with } S(\gamma, R) \leq R(\alpha + \varepsilon) \text{ and } \gamma(R) \in [0, R(\alpha + \varepsilon)]\} \cap \\ &\{\nexists \gamma \in \Gamma^R \text{ with } S(\gamma, R) \leq R(\alpha + \varepsilon) \text{ and } \gamma(R) \in [-R(\alpha + \varepsilon), 0]\} \end{aligned}$$



has probability less than  $\varepsilon^6$ . These two events are increasing functions of the Poisson processes and, by symmetry, have equal probabilities, so by the FKG inequalities (as in [BG]) we have

$$P[\nexists \gamma \in \Gamma^R \text{ with } S(\gamma, R) \leq R(\alpha + \varepsilon) \text{ and } \gamma(R) \in [0, R(\alpha + \varepsilon)]] < \varepsilon^3$$

, that is,

$$P[\exists \gamma \in \Gamma^R \text{ with } S(\gamma, R) \leq R(\alpha + \varepsilon) \text{ and } \gamma(R) \in [0, R(\alpha + \varepsilon)]] > 1 - \varepsilon^3$$

and, by symmetry,

$$P[\exists \gamma \in \Gamma^R \text{ with } S(\gamma, R) \leq R(\alpha + \varepsilon) \text{ and } \gamma(R) \in [-R(\alpha + \varepsilon), 0]] > 1 - \varepsilon^3$$

We remark that such paths must be contained in space time rectangle  $[-R(\alpha + \varepsilon), R(\alpha + \varepsilon)] \times [0, R]$ .

Thus outside probability strictly less than  $\frac{1}{\varepsilon}\varepsilon^3 = \varepsilon^2$ , we can "navigate" a path  $\gamma \in \Gamma^{\frac{R}{\varepsilon}}$  with  $S(\gamma, \frac{R}{\varepsilon}) \leq \frac{1}{\varepsilon}R(\alpha + \varepsilon)$ , which lies entirely in spacetime rectangle  $[-2R(\alpha + \varepsilon), 2R(\alpha + \varepsilon)] \times [0, \frac{R}{\varepsilon}]$  and which has  $\gamma(\frac{R}{\varepsilon}) \in [-R(\alpha + \varepsilon), R(\alpha + \varepsilon)]$ . Therefore we have with probability at least  $1 - \varepsilon^2$

there is a path  $\gamma \in \Gamma^{\frac{R}{\varepsilon}}$  so that (i)  $S(\gamma, \frac{R}{\varepsilon}) \leq R(\alpha + \varepsilon)(1 + 2\varepsilon)/\varepsilon$

(ii)  $\gamma$  lies entirely within  $[-2R(\alpha + \varepsilon), 2R(\alpha + \varepsilon)] \times [0, \frac{R}{\varepsilon}]$ .

Now provided that  $\delta$  is chosen sufficiently small we have also that with probability  $> 1 - \varepsilon^2$  we have  $\gamma$  satisfying in addition to (i) and (ii) above

(iii) No two jump times of  $\gamma$  are within  $2\delta$  of each other or of time 0 or time  $\frac{R}{\varepsilon}$ . Also the path  $\gamma$  is at all times at least  $2\delta$  away from points of  $N$  (considered now as a random subset of space time).

We define a 2-dependent oriented percolation scheme on  $\{(m, n) : n \geq 0, m + n \equiv 0(\text{mod}(2))\}$  as follows: We say that the bond from  $(m, n)$  to  $(m \pm 1, n + 1)$  is open if there is a path  $\gamma$  from  $(mR(\alpha + \varepsilon), n\frac{R}{\varepsilon})$  to  $((m \pm 1)R(\alpha + \varepsilon), (n + 1)\frac{R}{\varepsilon})$  that satisfies (i) and

(ii')  $\gamma$  lies entirely within  $[(m - 2)R(\alpha + \varepsilon), (m + 2)R(\alpha + \varepsilon)] \times [n\frac{R}{\varepsilon}, (n + 1)\frac{R}{\varepsilon}]$ .

(iii') No two jump times of  $\gamma$  are within  $2\delta$  of each other or of time  $n\frac{R}{\varepsilon}$  or time  $(n + 1)\frac{R}{\varepsilon}$ . Also the path  $\gamma$  is at all times at least  $2\delta$  away from points of  $N$

Then we have that (provided  $\varepsilon$  was chosen sufficiently small) the percolation system is supercritical (see the appendix of [D2], which while formally treating oriented bond percolation, is valid for our bond percolation). That is with probability one there is a point  $(0, n)$  with infinitely many "descendants".

**Lemma 3.1** *If  $k$  is sufficiently small then for all  $(m, n)$  if the percolation bond  $(m, n) \rightarrow (m \pm 1, n + 1)$  is open then with probability at least  $k^{\frac{R}{\varepsilon}(\alpha + \varepsilon)(1 + 3\varepsilon)}$  a random walk started at  $mR(\alpha + \varepsilon)$  at time  $n\frac{R}{\varepsilon}$  will survive until time  $(n + 1)\frac{R}{\varepsilon}$  and will be in position  $(m \pm 1)R(\alpha + \varepsilon)$  at this time.*

*Proof* Let a path satisfying (i), (ii') and (iii') be  $\gamma$ . Let its jumps be at times  $0 < t_1, t_2, \dots, t_r \leq R(\alpha + \varepsilon)(1 + 2\varepsilon)/\varepsilon$ . We consider the event that our random walk makes precisely  $r$  jumps in the time interval, these jumps occurring within the intervals  $(t_i - \delta/3, t_i + \delta/3)$  (one jump in each

interval) and the jumps are equal to the corresponding jumps of  $\gamma$ . This event is contained in the event of interest and has probability at least

$$e^{-\frac{R}{\varepsilon}k} \prod_{j=1}^r \left(\frac{2\delta}{3} \frac{k}{2}\right).$$

This is easily seen to exceed  $k^{\frac{R}{\varepsilon}(\alpha+\varepsilon)(1+3\varepsilon)}$  for  $k$  small.  $\square$

**Corollary 3.1**  $\frac{\lambda(k,d)}{\ln(\frac{1}{k})} \leq \alpha(d)$ .

*Proof* Given our percolation scheme we have (provided  $\varepsilon$  was chosen sufficiently small) that there exists  $n_0$  so that  $(0, n_0)$  is a point of percolation. That is to say there exists  $0 = m_0, m_1, \dots, m_j \dots$  so that  $\forall j \geq 1$ , the bond between  $(m_{j-1}, n_0 + j - 1)$  and  $(m_j, n_0 + j)$  is open.

It follows from induction and Lemma 3.1 that a random walk starting at site 0 at time  $n_0$  has chance at least  $k^{\frac{R}{\varepsilon}(\alpha+\varepsilon)(1+3\varepsilon)j}$  of surviving until time  $(n_0 + j)\frac{R}{\varepsilon}$  and being at  $m_j$  at this time. The chance that a random walk starting at site 0 at time 0 reaches site 0 at time  $n_0\frac{R}{\varepsilon}$  is strictly positive ( $\underline{N}$  a.s.). So we have for some  $c_k(\omega) > 0$  that

$$p\left(\left(n_0 + j\right)\frac{R}{\varepsilon}, N\right) \geq c_k(\omega)k^{\frac{R}{\varepsilon}(\alpha+\varepsilon)(1+3\varepsilon)j}$$

for  $k \leq k_0$ . Thus  $\lambda(k, d) \leq \ln(\frac{1}{k})(\alpha + \varepsilon)(1 + 3\varepsilon)$ . The corollary follows from the arbitrariness of  $\varepsilon$ .  $\square$

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